

Probability Theory 2 : Solution Sheet 7

Exercice 1

Let $\{t_n\}_{n \in \mathbb{N}^*}$ an enumeration of $[0, \infty) \cap \mathbb{Q}$. Set $D_n = \{t_1, \dots, t_n\}$ and $D = \bigcup_{n \in \mathbb{N}^*} D_n$.

1. By Doob inequality on D_n , we have

$$\lambda^p \mathbb{P} \left\{ \sup_{1 \leq k \leq n} |X_{t_k}| \geq \lambda \right\} \leq \sup_{1 \leq k \leq n} \mathbb{E}[|X_{t_k}|^p].$$

Since

$$\sup_{1 \leq k \leq n} |X_{t_k}| \nearrow \sup_{t \in D} |X_t| \quad \text{a.s.}$$

using monotone convergence theorem and

$$\sup_{1 \leq k \leq n} \mathbb{E}[|X_{t_k}|^p] \nearrow \sup_{t \in D} \mathbb{E}[|X_t|^p],$$

gives

$$\lambda^p \mathbb{P} \left\{ \sup_{t \in D} |X_t| \geq \lambda \right\} \leq \sup_{t \in D} \mathbb{E}[|X_t|^p].$$

Since (X_t) is right continuous, we have

$$\sup_{t \in D} |X_t| = \sup_{t \geq 0} |X_t|,$$

what prove the claim.

2. The proof is similar and is left to the reader.

Exercice 2

1. Let X a \mathcal{A} -measurable r.v. We have to prove that

$$\mathbb{E}[e^{t_1 X + t_2 Y}] = \mathbb{E}[e^{t_1 X}] \mathbb{E}[e^{t_2 Y}].$$

But this is rather clear since

$$\mathbb{E}[e^{t_1 X} e^{t_2 Y}] = \mathbb{E}[e^{t_1 X} \mathbb{E}[e^{t_2 Y} | \mathcal{A}]] = \mathbb{E}[e^{t_1 X}] \mathbb{E}[e^{t_2 Y}],$$

where the first equality come from the fact that $e^{t_1 X}$ is \mathcal{A} -measurable.

2. \Rightarrow Suppose (M_t^α) is a Martingale for all $\alpha \in \mathbb{R}$. Let $0 \leq s < t$. Then

$$\begin{aligned} \mathbb{E}[e^{\alpha(X_t - X_s)}] &= \mathbb{E} \left[e^{\alpha(X_t - X_s) - \frac{\alpha^2}{2}(t-s)} \right] \\ &= e^{\frac{\alpha^2}{2}(t-s)} \mathbb{E} \left[e^{\alpha(X_t - X_s) - \frac{\alpha^2}{2}(t-s)} \right] \\ &\stackrel{(1)}{=} e^{\frac{\alpha^2}{2}(t-s)} \mathbb{E} \left[e^{-\alpha X_s + \frac{\alpha^2}{2}s} \mathbb{E} \left[e^{\alpha X_t - \frac{\alpha^2}{2}t} | \mathcal{F}_s \right] \right] \\ &\stackrel{(2)}{=} e^{\frac{\alpha^2}{2}(t-s)} \end{aligned}$$

where we used the fact that $e^{\alpha X_s + \frac{\alpha^2}{2}s}$ is \mathcal{F}_s measurable in (1), and the martingale property in (2). Therefore $X_t - X_s \sim \mathcal{N}(0, t-s)$. Moreover, for all $\alpha \in \mathbb{R}$,

$$\mathbb{E}[e^{\alpha(X_t - X_s)} | \mathcal{F}_s] = e^{-\alpha X_s + \frac{\alpha^2}{2}s} \mathbb{E} \left[e^{\alpha X_t - \frac{\alpha^2}{2}t} | \mathcal{F}_s \right] = e^{\frac{\alpha^2}{2}(t-s)} = \mathbb{E}[e^{\alpha(X_t - X_s)}],$$

and thus, by question 1., $X_t - X_s$ and \mathcal{F}_s are independents. Since $X_0 = 0$ a.s. and is continuous, we conclude that (X_t) is a Brownian motion.

⊖ Suppose (X_t) is a Brownian motion. The fact that $\mathbb{E} \left[\left| e^{\alpha X_t - \frac{\alpha^2}{2} t} \right| \right] < \infty$ for all $t \geq 0$ and that $e^{\alpha X_t - \frac{\alpha^2}{2} t}$ is $\mathcal{F}_t := \sigma(X_s \mid s \leq t)$ adapted for all $t \geq 0$ are clear. Let $\alpha \in \mathbb{R}$ and $0 \leq s < t$.

$$\begin{aligned} \mathbb{E} \left[e^{\alpha X_t - \frac{\alpha^2}{2} t} \mid \mathcal{F}_s \right] &\stackrel{(3)}{=} e^{\alpha X_s - \frac{\alpha^2}{2} s} \mathbb{E} \left[e^{\alpha(X_t - X_s)} \mid \mathcal{F}_s \right] \\ &\stackrel{(4)}{=} e^{\alpha X_s - \frac{\alpha^2}{2} s} \mathbb{E} \left[e^{\alpha(X_t - X_s)} \right] \\ &\stackrel{(5)}{=} e^{\alpha X_s - \frac{\alpha^2}{2} s}, \end{aligned}$$

where we used the fact that $e^{\alpha X_s}$ is \mathcal{F}_s measurable in (3), that $X_t - X_s$ is independent of \mathcal{F}_s in (4) and that $X_t - X_s \sim \mathcal{N}(0, t - s)$ in (5). Therefore, $\left(e^{\alpha X_t - \frac{\alpha^2}{2} t} \right)$ is a martingale for all $\alpha \in \mathbb{R}$.

Exercise 3

1. Using reflexion principle yields

$$\begin{aligned} \mathbb{P}\{\tau \geq T\} &= \mathbb{P} \left\{ \sup_{t \in [0, T]} B_t \leq a \right\} \\ &= 1 - 2\mathbb{P}\{B_T > a\} \\ &= 1 - \frac{2}{\sqrt{2\pi T}} \int_{-\infty}^{-a} e^{-\frac{x^2}{2T}} dx \\ &= \frac{1}{\sqrt{2\pi T}} \int_{-a}^a e^{-\frac{x^2}{2T}} dx. \end{aligned}$$

Since

$$\int_{-a}^a e^{-\frac{x^2}{2T}} dx = \frac{1}{2} \int_0^a e^{-\frac{x^2}{2T}} dx \geq \frac{a}{2} e^{-\frac{a^2}{2T}},$$

we get

$$\mathbb{E}[\tau] = \int_0^\infty \mathbb{P}\{\tau \geq T\} dT = \infty.$$

2. Set

$$M_t = e^{\sqrt{2\lambda} B_t - \lambda t}.$$

By the previous exercise, (M_t) is a martingale. Then so is $(M_{\tau \wedge n})_{n \in \mathbb{N}}$. By the optional stopping theorem,

$$\mathbb{E}[M_{\tau \wedge n}] = \mathbb{E}[M_0] = 1.$$

Since

$$|M_{\tau \wedge n}| = e^{\sqrt{2\lambda} B_{\tau \wedge n} - \lambda(\tau \wedge n)} \leq e^{\sqrt{2\lambda} a},$$

we have by DCT (dominated convergence theorem) that

$$\mathbb{E}[M_\tau] = \lim_{n \rightarrow \infty} \mathbb{E}[M_{\tau \wedge n}] = 1.$$

Therefore

$$\mathbb{E}[e^{\sqrt{2\lambda} a - \lambda \tau}] = 1,$$

and thus

$$\mathbb{E}[e^{-\lambda \tau}] = e^{-2\lambda a}.$$